

ML & Opt lecture 4

- Recall conditional distributions as capturing all uncertainties
- Point estimates
- Expectations & Variances revisited
- Conditional Expectations
- Regression Functions — best point estimate for least squares regression

Conditional Distributions

Given r.v.s Y and X ,

$$p(y|x) = \frac{p(x, y)}{p(x)}$$

captures the dependence of y on x

$$1) \quad p(y, z|x) = p(y|x) p(z|x)$$

conditional independence $(y \perp\!\!\!\perp z | x)$

$$\Leftrightarrow p(y|z, x) = p(y|x)$$

Show in participation
1

$$2) \quad p(x|z) = \int p(x, y|z) dy$$

marginalizing joint
conditional gives
marginal conditional

Conditional distributions capture everything we know about y given x .

But often we want an informative single number describing the behavior of y , conditioned on x .

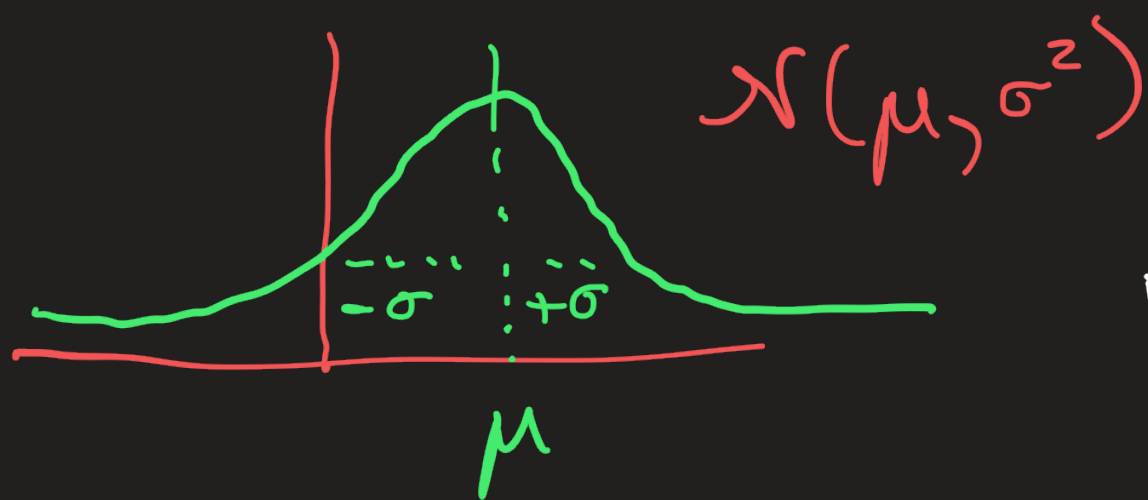
Call these point estimates.

Expected Values (point estimates of R.V.s)

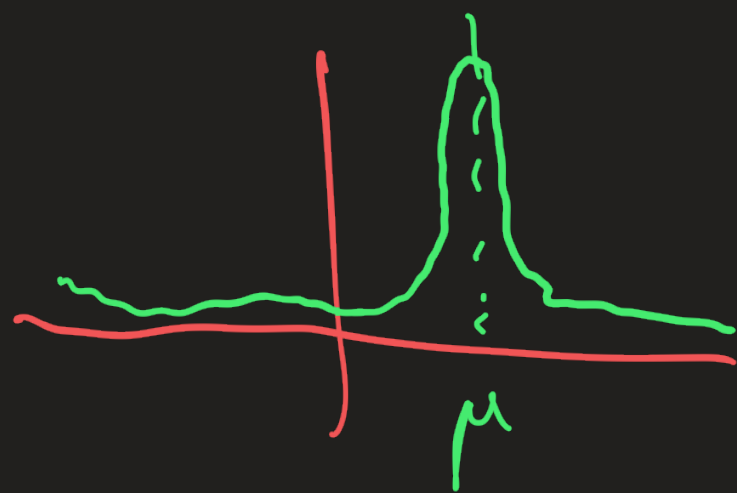
$$\mathbb{E}X = \sum_x x p(X=x)$$

weighted average of
the possible values a r.v.
takes on

$$= \int x p(x) dx$$



$$\mathbb{E}X = \int_{-\infty}^{\infty} x p(x) dx$$

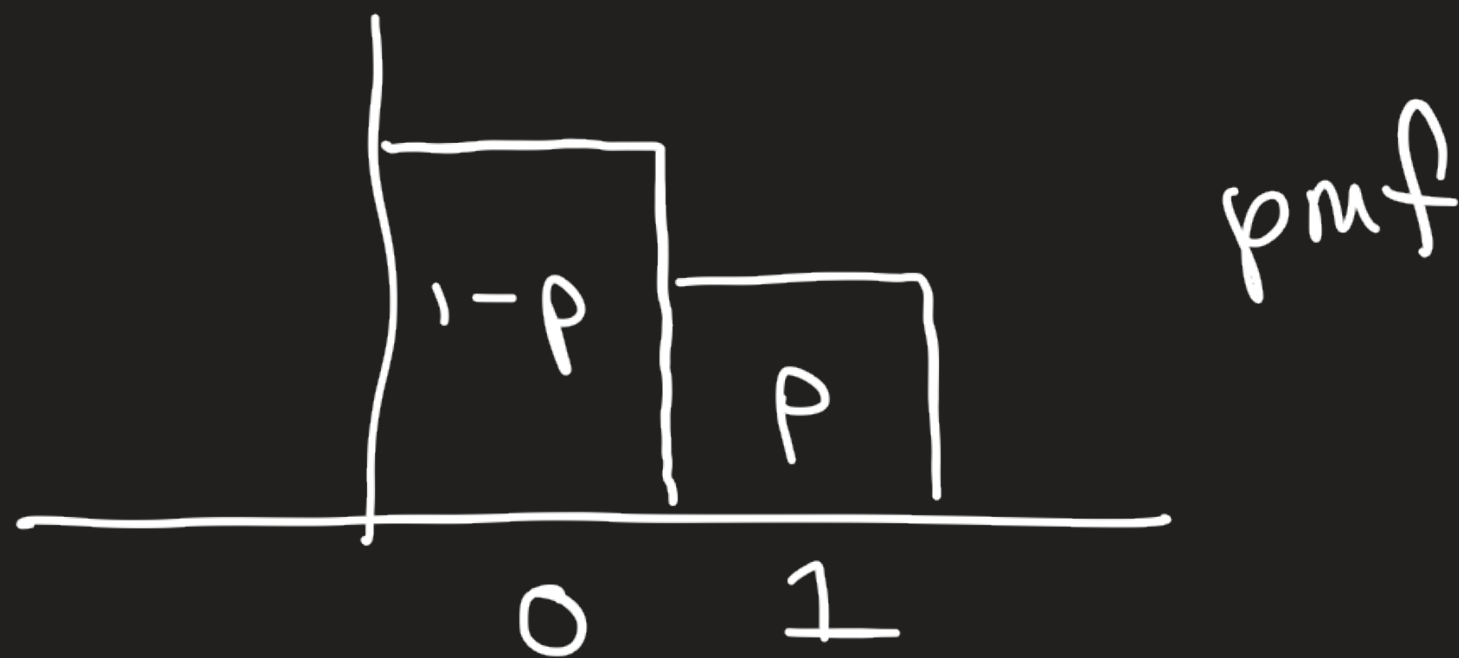


$$= \int_{-\infty}^{\infty} x \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right) dx$$
$$= \mu$$

Bernoulli r.v.s

$$X = \begin{cases} 0 & 1-p \\ 1 & p \end{cases}$$

where $0 \leq p \leq 1$



$$\mathbb{E} X = 0 \cdot (1-p) + 1 \cdot p$$

Variance : how well does $\mathbb{E}X$ describe X , on average?

$$\text{Var}(X) = \mathbb{E}(X - \mathbb{E}X)^2$$

$$= \mathbb{E}[X^2 - 2X\mathbb{E}X + (\mathbb{E}X)^2]$$

$$= \mathbb{E}[X^2] - 2\mathbb{E}[X\mathbb{E}X] + (\mathbb{E}X)^2$$

$$= \mathbb{E}[X^2] - 2(\mathbb{E}X)^2 + (\mathbb{E}X)^2$$

$$= \mathbb{E}[X^2] - (\mathbb{E}X)^2$$

↑
2nd moment

↑
square of the 1st moment

$$\int x \mathbb{E}X p(x) dx$$

$$\mathbb{E}X \int x p(x) dx$$

$$(\mathbb{E}X)^2$$

$X \sim \mathcal{N}(\mu, \sigma^2)$ then

$$\text{Var}(X) = \mathbb{E}(X - \mathbb{E}X)^2 = \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} (x - \mu)^2 \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right) dx$$

$$= \mathbb{E}[X^2] - (\mathbb{E}X)^2$$

$$= \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} x^2 \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right) dx - \mu^2$$

$$= \dots = \sigma^2$$

$$X \sim \text{Bern}(p)$$

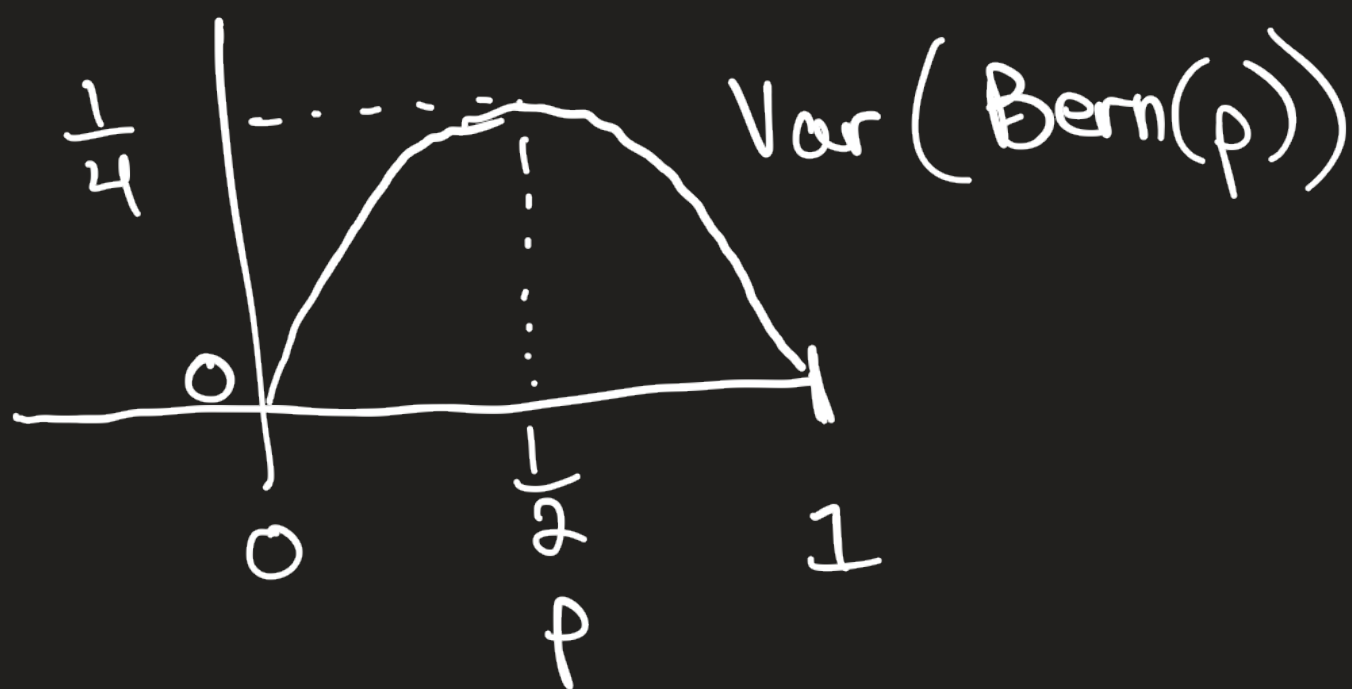
$$\mathbb{E}X = p$$

because $X = X^2$ since X is 0 or 1

$$\text{Var}(X) = \mathbb{E}(X^2) - (\mathbb{E}X)^2 \stackrel{\swarrow}{=} \mathbb{E}X - (\mathbb{E}X)^2$$

$$= p(1-p)$$

so if $p \approx 1$ or $p \approx 0$ $\text{Var}(X) \approx 0$



Expectations & Variances & Independence

X & Y are independent then

$$\mathbb{E}(XY) = (\mathbb{E}X)(\mathbb{E}Y)$$

$$\mathbb{E}XY = \int xy p(x, y) dx dy$$

$$\text{by } X \perp Y = \int xy p(x) p(y) dx dy$$

$$= \int x p(x) dx \int y p(y) dy$$

$$= (\mathbb{E}X)(\mathbb{E}Y)$$

If X & Y are independent

$$\text{Var}(\alpha X + Y) = \alpha^2 \text{Var}(X) + \text{Var}(Y)$$

$$\text{Var}(\alpha X + Y) = \mathbb{E}(\alpha X + Y)^2 - [\mathbb{E}(\alpha X + Y)]^2$$

$$= \mathbb{E}(\alpha X + Y)^2 - [\alpha \mathbb{E}X + \mathbb{E}Y]^2$$

$$= \mathbb{E}(\alpha^2 X^2 + 2\alpha XY + Y^2) - [\alpha^2 (\mathbb{E}X)^2 + 2\alpha \mathbb{E}X \mathbb{E}Y + (\mathbb{E}Y)^2]$$

○ because $X \perp Y$
" " " " " "

$$= \alpha^2 \mathbb{E}\{X^2\} + 2\alpha \mathbb{E}\{XY\} + \mathbb{E}\{Y^2\} - \alpha^2 (\mathbb{E}X)^2 - 2\alpha \mathbb{E}X \mathbb{E}Y - (\mathbb{E}Y)^2$$
$$= \alpha^2 \text{Var}(X) + \text{Var}(Y)$$

Q: if X & Y are not independent
what can we say about $\mathbb{E}[XY]$
and $\text{Var}(X + Y)$?

A: not anything useful, in general

e.g. $Y = -X$ then $\text{Var}(X + Y) = \text{Var}(0) = 0$

and if $Y = \frac{1}{X}$ then $\mathbb{E}[XY] = 1 \neq \mathbb{E}(X)\mathbb{E}(\frac{1}{X})$

Consequence of Variance & Independence

Sometimes we want to estimate $\mathbb{E}X$ by sampling n independent copies of X , $X_1, \dots, X_n$ and claiming

$$\mathbb{E}X \approx \frac{1}{n} \sum_{i=1}^n X_i = S_n$$

Law of Large Numbers

Intuition for why this is true:

$$\mathbb{E}S_n = \frac{1}{n} \sum_{i=1}^n \mathbb{E}X_i = \frac{1}{n} n \mathbb{E}X = \mathbb{E}X$$

and

$$\text{Var}(S_n) = \text{Var}\left(\frac{1}{n}X_1 + \dots + \frac{1}{n}X_n\right)$$

$$\begin{aligned} &= \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i) = \frac{1}{n^2} n \text{Var}(X) \\ &= \frac{1}{n} \text{Var}(X) \end{aligned}$$

$$\mathbb{E} S_n = \mathbb{E} \left(\frac{1}{n} \sum_{i=1}^n x_i \right) = \int_{\mathbb{R}^n} \left(\frac{1}{n} \sum_{i=1}^n x_i \right) p(x_1, \dots, x_n) dx$$

$$= \int_{\mathbb{R}^n} \left(\frac{1}{n} \sum_{i=1}^n x_i \right) p(x_1) \dots p(x_n) dx$$

$$= \frac{1}{n} \sum_{i=1}^n \int_{\mathbb{R}^n} x_i p(x_1) \dots p(x_n) dx$$

$$= \frac{1}{n} \sum_{i=1}^n \int_{\mathbb{R}} x_i p(x_i) dx_i = \frac{1}{n} n \mathbb{E} X = \mathbb{E} X$$

Recall Population Risk Minimization vs ERM

$$\text{PRM: } f^* = \underset{f \in \mathcal{F}}{\operatorname{argmin}} \mathbb{E} \ell(f(x), y)$$

but in general we don't know $p(x, y)$

so sample (i.i.d.) copies of data (x_i, y_i)

and we approximate

$$\mathbb{E} \ell(f(x), y) \approx \frac{1}{n} \sum_{i=1}^n \underline{\ell(f(x_i), y_i)}$$

and take

$$\text{ERM: } \hat{f} = \underset{f \in \mathcal{F}}{\operatorname{argmin}} \frac{1}{n} \sum_{i=1}^n \ell(f(x_i), y_i)$$

Conditional Expectation

point estimate for the r.v. $y | X=x$

$$\mathbb{E}[y | X=x] = \int y p(y|x) dy$$

Example

		y		
		0	1	2
x	0	$\frac{2}{6}$	$\frac{1}{12}$	$\frac{1}{12}$
	1	$\frac{1}{12}$	$\frac{2}{6}$	$\frac{1}{12}$

$p(y x=0)$		
0	1	2
$\frac{2}{3}$	$\frac{1}{6}$	$\frac{1}{6}$

$p(y x=1)$		
0	1	2
$\frac{1}{6}$	$\frac{2}{3}$	$\frac{1}{6}$

$$\mathbb{E}[y | X=0] = 0 \cdot \frac{2}{3} + 1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} = \frac{1}{2}$$

$$\mathbb{E}[y | X=1] = 0 \cdot \frac{1}{6} + 1 \cdot \frac{2}{3} + 2 \cdot \frac{1}{6} = 1$$

$$\begin{aligned}\mathbb{E} Y &= 0 \cdot \left(\frac{2}{6} + \frac{1}{12}\right) + 1 \cdot \left(\frac{1}{12} + \frac{2}{6}\right) + 2 \cdot \left(\frac{1}{12} + \frac{1}{12}\right) \\ &= \frac{5}{12} + \frac{1}{3} = \frac{5}{12} + \frac{4}{12} = \frac{9}{12} = \frac{3}{4}\end{aligned}$$

By Law of Total Expectation

$$\mathbb{E}[\mathbb{E}[Y|X]] = \mathbb{E} Y$$

Check:

$$\begin{aligned}\mathbb{E}[\mathbb{E}[Y|X]] &= \mathbb{E}[Y|X=0] p(X=0) \\ &\quad + \mathbb{E}[Y|X=1] p(X=1) \\ &= \frac{1}{2} \cdot \frac{1}{2} + 1 \cdot \frac{1}{2} = \frac{1}{4} + \frac{1}{2} = \frac{3}{4} = \mathbb{E} Y\end{aligned}$$

Notice that $\mathbb{E}[Y|X]$ is a random variable
and $\mathbb{E}[Y|X=x]$ is a number

Properties

I) if y is a function of x , $y=f(x)$
then $\mathbb{E}[Y|X] = y$

II) if $y \perp x$, $\mathbb{E}[Y|X] = \mathbb{E}[Y]$ b/c $y \perp x$

$$\begin{aligned}\mathbb{E}[Y|X=x] &= \int y p(y|x) dy = \int y p(y) dy \\ &= \mathbb{E}Y\end{aligned}$$

III) For any function f

$$\mathbb{E}[y f(x) | x] = f(x) \mathbb{E}[y | x]$$

IV) Linearity

$$\mathbb{E}[\alpha y + z | x] = \alpha \mathbb{E}[y | x] + \mathbb{E}[z | x]$$

V) Total Expectation

$$\mathbb{E}[\mathbb{E}[y | x]] = \mathbb{E}[y]$$

Point Estimates for least squares regression

$p(y|x)$ tells us everything about y given x
but we want a point estimate $y = f(x)$

$$f^* = \underset{f}{\operatorname{argmin}} \mathbb{E} (y - f(x))^2$$

Note by the law of total expectation

$$\mathbb{E} (y - f(x))^2 = \mathbb{E} \left[\mathbb{E} [(y - f(x))^2 | x = x] \right]$$

so f^* can be obtained by choosing

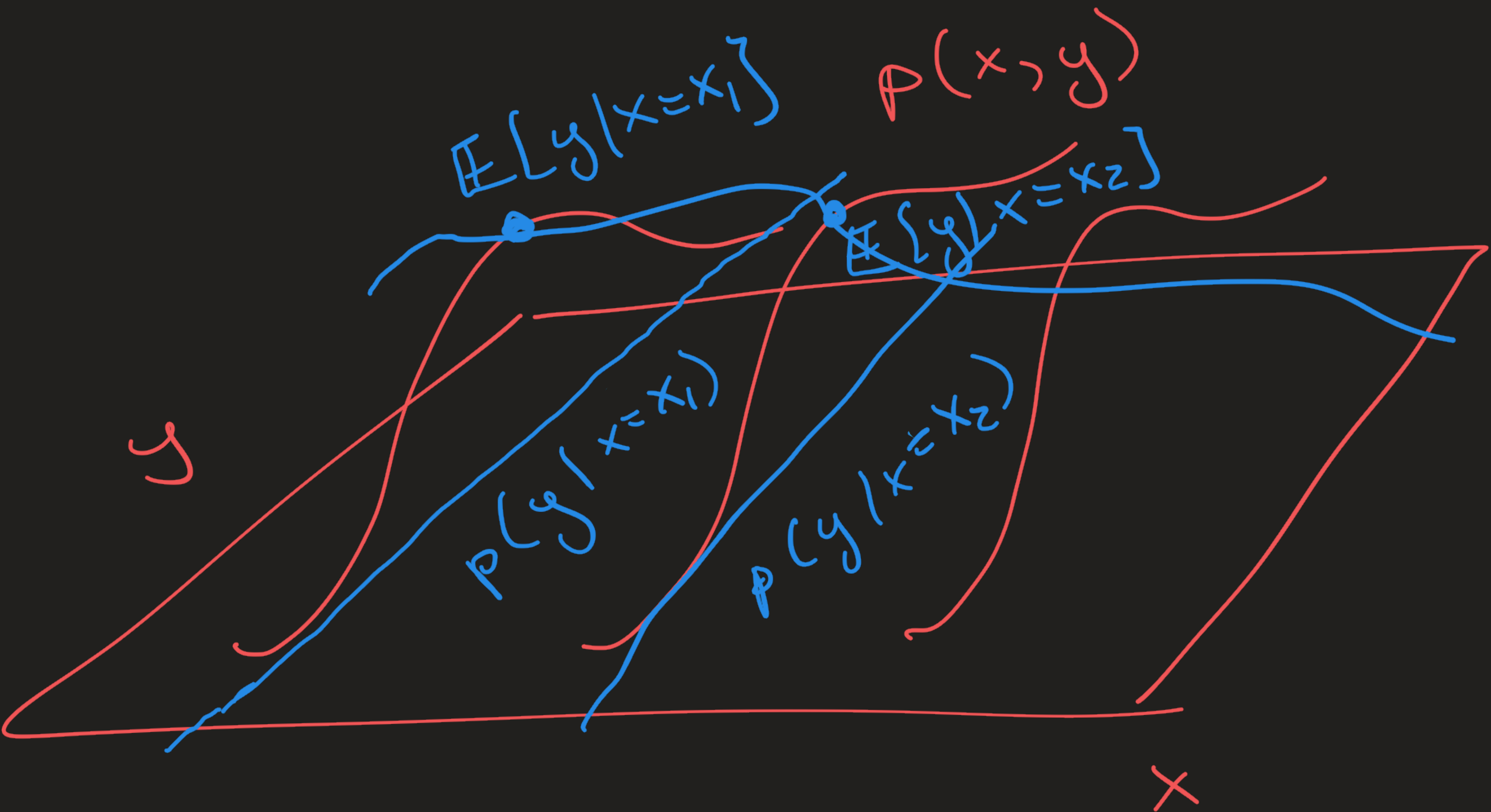
$$f^*(x) = \underset{c}{\operatorname{argmin}} \mathbb{E} [(y - c)^2 | x = x]$$

Notice that this is the BEST possible estimate for y , given x , in the least squares sense

Call

$$\eta(x) = \mathbb{E}[y|x]$$

the regression function



$$f(x) = \mathbb{E}[y | x=x]$$

